

Supplementary Material

Kernel-based sensitivity indices for any model behavior and screening

A: Proof of Proposition 1

Knowing that the density of \mathbf{X} is given by $\rho(\mathbf{x}) = c(F_1(x_1), \dots, F_d(x_d)) \prod_{j=1}^d \rho_j(x_j)$, we can write thanks to Equation (3)

$$\rho^w(\mathbf{x}) = \frac{w(\mathbf{x})}{\mathbb{E}_{F[w(\mathbf{X})]}} c(F_1(x_1), \dots, F_d(x_d)) \prod_{j=1}^d \rho_j(x_j).$$

B: Proof of Theorem 1

The following proof is simpler than the general one provided in Lamboni (2023). Since $\rho^w(\mathbf{x}) = \frac{w_e(\mathbf{x})}{\mathbb{E}_{F_{ind}[w_e(\mathbf{Y})]}} \prod_{j=1}^d \rho_j(x_j)$ (see Proposition 1), the density of $\mathbf{X}_{\sim u}^w \mid \mathbf{X}_u^w$ becomes

$$\rho_{\sim u \mid u}^w(\mathbf{x}_{\sim u} \mid \mathbf{x}_u^w) = \frac{w_e(\mathbf{x}_u^w, \mathbf{x}_{\sim u})}{\mathbb{E}_{F_{ind}[w_e(\mathbf{x}_u^w, \mathbf{Y}_{\sim u})]}} \prod_{j \in (\sim u)} \rho_j(x_j),$$

and we can write (bearing in mind $(\sim u) = (\pi_1, \dots, \pi_{|\pi|})$)

$$F_{\sim u \mid u}^w(\mathbf{x}_{\sim u} \mid \mathbf{x}_u^w) = \mathbb{E}_{F_{ind}} \left[\frac{w_e(\mathbf{x}_u^w, \mathbf{Y}_{\sim u})}{\mathbb{E}_{F_{ind}[w_e(\mathbf{x}_u^w, \mathbf{Y}_{\sim u})]}} \prod_{j=1}^{|\pi|} \mathbb{1}_{[-\infty, x_{\pi_j}]}(Y_{\pi_j}) \right].$$

Knowing that $X_{\pi_j} \stackrel{d}{=} F_{\pi_j}^{-1}(U_{\pi_j})$ with $U_{\pi_j} \sim \mathcal{U}(0,1)$ and using the theorem of transfer, we have

$$F_{\sim u \mid u}^w(\mathbf{x}_{\sim u} \mid \mathbf{x}_u^w) = \mathbb{E}_{\mathbf{U}_{\pi}} \left[\frac{w_e(\mathbf{x}_u^w, F_{\pi_1}^{-1}(U_{\pi_1}), \dots, F_{\pi_{|\pi|}}^{-1}(U_{\pi_{|\pi|}}))}{\mathbb{E}_{F_{ind}[w_e(\mathbf{x}_u^w, \mathbf{Y}_{\sim u})]}} \prod_{j=1}^{|\pi|} \mathbb{1}_{[-\infty, x_{\pi_j}]}(F_{\pi_j}^{-1}(U_{\pi_j})) \right],$$

with $\mathbf{U}_{\pi} := (U_{\pi_j}, j = 1, \dots, |\pi|) \sim \mathcal{U}(0,1)^d$. As F_j is strictly increasing, we have

$$\begin{aligned} F_{\sim u \mid u}^w(\mathbf{x}_{\sim u} \mid \mathbf{x}_u^w) &= \mathbb{E}_{\mathbf{U}_{\pi}} \left[\frac{w_e(\mathbf{x}_u^w, F_{\pi_1}^{-1}(U_{\pi_1}), \dots, F_{\pi_{|\pi|}}^{-1}(U_{\pi_{|\pi|}}))}{\mathbb{E}_{F_{ind}[w_e(\mathbf{x}_u^w, \mathbf{Y}_{\sim u})]}} \prod_{j=1}^{|\pi|} \mathbb{1}_{[0, F_{\pi_j}(x_{\pi_j})]}(U_{\pi_j}) \right] \\ &= \int_0^{F_{\pi_1}(x_{\pi_1})} \dots \int_0^{F_{\pi_{|\pi|}}(x_{\pi_{|\pi|}})} \frac{w_e(\mathbf{x}_u^w, F_{\pi_1}^{-1}(v_{\pi_1}), \dots, F_{\pi_{|\pi|}}^{-1}(v_{\pi_{|\pi|}}))}{\mathbb{E}_{F_{ind}[w_e(\mathbf{x}_u^w, \mathbf{Y}_{\sim u})]}} \prod_{j=1}^{|\pi|} dv_{\pi_j}. \end{aligned}$$

Now, if we use $\mathbf{V}_{\pi} := (V_{\pi_k} \sim \mathcal{U}(0, u_{\pi_k}), k = 1, \dots, |\pi|)$ for a random vector of independent variables and

$$\begin{aligned}
W(\mathbf{u}_\pi; \mathbf{x}_u^w) &:= \int_0^{u_{\pi_1}} \dots \int_0^{u_{\pi_{|\pi|}}} \frac{w_e \left(\mathbf{x}_u^w, F_{\pi_1}^{-1}(v_{\pi_1}), \dots, F_{\pi_{|\pi|}}^{-1}(v_{\pi_{|\pi|}}) \right)}{\mathbb{E}_{F_{ind}} [w_e(\mathbf{x}_u^w, \mathbf{Y}_{\sim u})]} \prod_{j=1}^{|\pi|} dv_{\pi_j} \\
&= \frac{\mathbb{E}_{\mathbf{V}_\pi} \left[w_e \left(\mathbf{x}_u^w, F_{\pi_1}^{-1}(V_{\pi_1}), \dots, F_{\pi_{|\pi|}}^{-1}(V_{\pi_{|\pi|}}) \right) \right]}{\mathbb{E}_{F_{ind}} [w_e(\mathbf{x}_u^w, \mathbf{Y}_{\sim u})]} \prod_{j=1}^{|\pi|} u_{\pi_j},
\end{aligned}$$

then W is a CDF of a random vector having $(0,1)^{d-|\mathbf{u}|}$ as the support, and we have

$$F_{\sim u|u}^w(\mathbf{x}_{\sim u} | \mathbf{x}_u^w) = W \left(F_{\pi_1}(x_{\pi_1}), \dots, F_{\pi_{|\pi|}}(x_{\pi_{|\pi|}}); \mathbf{x}_u^w \right).$$

C: Proof of Lemma 1

Since $\mathbf{X}_u^w, \mathbf{X}_u^{w'}$ are i.i.d. and $\Theta = \{\theta_0\}$, we can write thanks to Proposition 2

$$\begin{aligned}
f_u^{fo}(\mathbf{X}_u^w) &= \mathbb{E}_U [f(\mathbf{X}_u^w, r(\mathbf{X}_u^w, \mathbf{U}))] - \frac{\mathbb{E}[f(\mathbf{Y}_u, r(\mathbf{Y}_u, \mathbf{U}))w_e(\mathbf{Y})]}{\mathbb{E}[w_e(\mathbf{Y})]} \\
&= \mathbb{E}_U \left[f(\mathbf{X}_u^w, r(\mathbf{X}_u^w, \mathbf{U})) - \frac{\mathbb{E}_Y [f(\mathbf{Y}_u, r(\mathbf{Y}_u, \mathbf{U}))w_e(\mathbf{Y})]}{\mathbb{E}[w_e(\mathbf{Y})]} \right] \\
&= \mathbb{E}_U [f_u^{tot}(\mathbf{X}_u^w, \mathbf{U})].
\end{aligned}$$

Using the convexity of ϕ and the definition of the kernel, the Jensen inequality yields

$$\mathbb{E} \left[k \left(f_u^{fo}(\mathbf{X}_u^w), f_u^{fo}(\mathbf{X}_u^{w'}) \right) \right] \leq \mathbb{E} \left[k \left(f_u^{tot}(\mathbf{X}_u^w, \mathbf{U}), f_u^{tot}(\mathbf{X}_u^{w'}, \mathbf{U}') \right) \right].$$

Now, we are going to show that the total index is less than one. Let us consider the Dirac probability measure

$$\delta_{\mathbf{U}}(\mathbf{U}') := \delta_{\mathbf{0}}(\mathbf{U}' - \mathbf{U}) \text{ and the zero-mean expression of the outputs } f_u^c(\mathbf{X}_u^w, \mathbf{U}) = \mathbb{E} \left[f(\mathbf{X}_u^w, r(\mathbf{X}_u^w, \mathbf{U})) - \frac{\mathbb{E}_Y [f(\mathbf{Y}_u, r(\mathbf{Y}_u, \mathbf{U}'))w_e(\mathbf{Y})]}{\mathbb{E}[w_e(\mathbf{Y})]} \mid \mathbf{U}, \mathbf{X}_u^w \right].$$

$$\begin{aligned}
&\mathbb{E}[f_u^c(\mathbf{X}_u^w, \mathbf{U}) \mid \delta_{\mathbf{U}}(\mathbf{U}'), \mathbf{U}, \mathbf{X}_u^w] \\
&= \mathbb{E} \left[\mathbb{E} \left[f(\mathbf{X}_u^w, r(\mathbf{X}_u^w, \mathbf{U})) - \frac{\mathbb{E}_Y [f(\mathbf{Y}_u, r(\mathbf{Y}_u, \mathbf{U}'))w_e(\mathbf{Y})]}{\mathbb{E}[w_e(\mathbf{Y})]} \mid \mathbf{U}, \mathbf{X}_u^w \right] \mid \delta_{\mathbf{U}}(\mathbf{U}'), \mathbf{U}, \mathbf{X}_u^w \right] \\
&= \mathbb{E} \left[\mathbb{E} \left[f(\mathbf{X}_u^w, r(\mathbf{X}_u^w, \mathbf{U})) - \frac{\mathbb{E}_Y [f(\mathbf{Y}_u, r(\mathbf{Y}_u, \mathbf{U}'))w_e(\mathbf{Y})]}{\mathbb{E}[w_e(\mathbf{Y})]} \mid \delta_{\mathbf{U}}(\mathbf{U}'), \mathbf{U}, \mathbf{X}_u^w \right] \mid \mathbf{U}, \mathbf{X}_u^w \right] \\
&= \mathbb{E} \left[f(\mathbf{X}_u^w, r(\mathbf{X}_u^w, \mathbf{U})) - \frac{\mathbb{E}_Y [f(\mathbf{Y}_u, r(\mathbf{Y}_u, \mathbf{U}))w_e(\mathbf{Y})]}{\mathbb{E}[w_e(\mathbf{Y})]} \mid \mathbf{U}, \mathbf{X}_u^w \right] = f_u^{tot}(\mathbf{X}_u^w, \mathbf{U}),
\end{aligned}$$

bearing in mind the formal definition of conditional expectation. The second result holds by applying the conditional Jensen inequality.

For the upper bound of the total index, knowing that $f_u^*(\mathbf{X}_u^w, \mathbf{X}_u^{w'}, \mathbf{U}) = f(\mathbf{X}_u^w, r(\mathbf{X}_u^w, \mathbf{U})) - f(\mathbf{X}_u^{w'}, r(\mathbf{X}_u^{w'}, \mathbf{U}))$, we can write $f_u^{tot}(\mathbf{X}_u^w, \mathbf{U}) = \mathbb{E}_{\mathbf{X}_u^{w'}} [f_u^*(\mathbf{X}_u^w, \mathbf{X}_u^{w'}, \mathbf{U})]$, and the result follows.

D: Proof of Theorem 2

First, the consistency of the estimators holds by applying the Slutsky theorem bearing in mind the Taylor expansion, that is,

$$k\left(\widehat{f}_u^{fo}(\mathbf{Y}_{i,u}), \widehat{f}_u^{fo}(\mathbf{Y}'_{i,u})\right) = k\left(f_u^{fo}(\mathbf{Y}_{i,u}), f_u^{fo}(\mathbf{Y}'_{i,u})\right) \\ + \nabla^T k\left(f_u^{fo}(\mathbf{Y}_u), f_u^{fo}(\mathbf{Y}'_u)\right) \begin{bmatrix} \widehat{f}_u^{fo}(\mathbf{Y}_{i,u}) - f_u^{fo}(\mathbf{Y}_{i,u}) \\ \widehat{f}_u^{fo}(\mathbf{Y}'_{i,u}) - f_u^{fo}(\mathbf{Y}'_{i,u}) \end{bmatrix} + R_{m_1},$$

with $R_{m_1} \xrightarrow{P} 0$ when $m_1 \rightarrow \infty$. We obtain the results by applying the law of large numbers. Second, the central limit theorem ensures that

$$\sqrt{m} \left(\frac{1}{m} \sum_{i=1}^m k\left(\widehat{f}_u^{fo}(\mathbf{Y}_{i,u}), \widehat{f}_u^{fo}(\mathbf{Y}'_{i,u})\right) w_e(\mathbf{Y}_i) w_e(\mathbf{Y}'_i) - D_u^k \right) \xrightarrow{D} \mathcal{N}(0, \sigma_u^{fo}),$$

with $D_u^k = \mathbb{E} \left[k\left(f_u^{fo}(\mathbf{Y}_u), f_u^{fo}(\mathbf{Y}'_u)\right) w_e(\mathbf{Y}) w_e(\mathbf{Y}') \right]$.

Third, the asymptotic distributions are straightforward using the Slutsky theorem under the technical assumption $m/M \rightarrow 0$, $m_1/M \rightarrow 0$ (see Lamboni (2020b, 2019) for more details).

E: Derivation of SFs used in Section 5.1

Using the model output and the dependency models, we can write

$$\begin{aligned} f_1^{fo}(X_1^w) &= (X_1^w)^2 (1 - \mathbb{E}[Z_2] - \mathbb{E}[Z_3(1 - Z_2)]) - \mathbb{E}[(X_1^w)^2] (1 - \mathbb{E}[Z_2] - \mathbb{E}[Z_3(1 - Z_2)]) \\ &= [(X_1^w)^2 - \mathbb{E}[(X_1^w)^2]] (1 - \mathbb{E}[Z_2] - \mathbb{E}[Z_3(1 - Z_2)]) \\ &= [(X_1^w)^2 - c/5] (1 - 1/4 - 1/3(1 - 1/4)) = \frac{1}{2} [(X_1^w)^2 - c/5]; \\ f_1^{tot}(X_1^w, Z_2, Z_3) &= [(X_1^w)^2 - \mathbb{E}[(X_1^w)^2]] (1 - Z_2 - Z_3(1 - Z_2)) = [(X_1^w)^2 - c/5] (1 - Z_2 - Z_3(1 - Z_2)). \end{aligned}$$

We also have

$$\begin{aligned} f^c(X_1^w, Z_2, Z_3) &= f(X^w) - c/10 - c/4 - c/4 = (X_1^w)^2 (1 - Z_2 - Z_3(1 - Z_2)) + cZ_2 + cZ_3(1 - Z_2) - \frac{3}{5}c. \\ f^*(X_1^w, X_1^{w'}, Z_2, Z_3) &= [(X_1^w)^2 - (X_1^{w'})^2] (1 - Z_2 - Z_3(1 - Z_2)). \end{aligned}$$